Financial Statement Comparability and New Debt Issues

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Abstract

Using a large sample of firms issuing new debts, this paper investigates how a firm's financial statements comparability affects its cost of new debt issues. We predict and find that higher comparability is associated with (1) higher bond ratings and (2) lower bond yield spreads when companies issue new debts. Our results are consistent with the view that bond rating analysts and bond investor favor greater comparability when they evaluate new bonds and make investment decisions.

Keywords: comparability, new debt issues, bond ratings, bond yield

1. Introduction

In this study, we examine the impact of accounting comparability on the cost of new debt issues. The Financial Accounting Standards Board's (FASB) Statement of Financial Accounting Concepts (SFAC) No. 8, Qualitative Characteristics of Useful Financial Information, defines comparability as "the qualitative characteristic that enables users to identify and understand similarities in, and differences among, financial statement items". Accounting comparability is arguably one of the most important accounting qualities (Kothari et al., 2010), and we complement emergent studies on the impact of accounting quality by investigating its effect on one significant third-party cost, namely, the cost of new debt issues.

We predict the higher accounting comparability a firm's financial reports have, the lower cost of debt the firm will incur when the firm goes to the capital market and issues new debts. Our hypothesis is based on theoretic research, as well as recent empirical evidence, which examines the relationship between information environment and cost of debt. Bharath, J. Sunder, and S. Sunder (2008) find that financial reporting quality can significantly reduce information asymmetry, improve the information environment for creditors to assess the creditworthiness of borrowers, and lower financing cost of a firm consequently. De Franco et al. (2011) empirically demonstrate that accounting comparability is positively related to information environment improvement and increases both the quantity and the quality of information available to outsiders of a firm to help them evaluate financial health and firm performance of the firm.

Our sample consists of 1,981 new corporate bond issues from 1990 to 2014. The bonds in our sample are issued by non-financial firms and backed solely by the issuers' ability to pay. The sample does not include convertible bonds, mortgage bonds, asset-backed bonds, or deferred interest bonds. This allows us to examine bonds which are based entirely on the issuing firm's default risk and not the risk of another underlying asset or option. The sample is restricted to new issues because this helps ensure the rating is the result of a detailed and complete recent analysis conducted by bond rating analysts and avoids problems related to any potential delay in changing existing ratings which have been expressed in the popular press (e.g., Wall Street Journal 2011, 2012). Moreover, because each issue is new, we do not have to worry about any timing delays in the reflection of current price due to bonds that are thinly traded. Therefore, new bond issues provide a good sample to investigate our question. Our results indicate that firms with higher accounting comparability receive higher bond ratings. Besides that, we also find the bond yields are lower for firms with higher accounting comparability, consistent with that bond investors require lower yield premium when the firms' accounting comparability is high.

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We contribute to the general literature on the impact of financial reporting quality on capital market participants (Easley and O'Hara 2004; Lambert, Leuz, and Verrecchia 2007). Prior literature suggests a positive influence of financial reporting quality on capital market participants, such as financial analysts and equity investors (De Franco et al., 2011; Kothari et al., 2010). We specifically complete studies of Fang, Li, Xin, and Zhang (2016), Kim, Kraft, and Ryan (2013), and Imhof, Seavey and Smith (2017) that suggest that comparability can reduce information asymmetry and information risk, and consequently cost of capital. Our evidence suggests that when firms issue new debts, accounting comparability is an important consideration in cost of debt.

Our research also has strong practical implications. Our empirical evidence indicates a substantial impact of financial reporting quality on cost of issuing new debts, which is a significant business cost. This economically significant cost saving may have nontrivial importance for any firms.

There is one caveat when interpreting the empirical results. In this paper, equity-market based comparability based on De Franco et al. (2011) is used. Kim et al. (2013) develop a debt-market based comparability measure based on Moody's' adjustments to firms' accounting numbers compiled in Financial Metrics database. We use equity-market based comparability measures proposed by De Franco et al (2011) due to the limitation in access of the Financial Metrics database. However, this less refined measure is expected to work against our findings.

Our study proceeds as follows. Section 2 reviews the related literature and develops the hypothesis. Section 3 discusses the research design while section 4 presents the results. Section 5 concludes the paper.

2. Literature Review and Hypothesis Development

2.1 Bond Ratings

Bond ratings are important as they provide an independent appraisal to the market regarding the default risk associated with a firm's debt. Ratings reflect the analysts' perception of a firm's probability of default on future payments. The calculation of this probability depends on a wide range of factors, including the profitability ratios of the firm (Standard &Poor's Ratings Group, 2008). Information which provides insight into the quality and potential persistence of a firm's earnings should be important to analysts in establishing an issue's rating.

From an economic perspective, firms raise substantially more "new" funds in the credit market every year than in the equity market to finance new and continuing activities. For example, Thomson Reuters reports that US offerings of debt were \$1.3 trillion in 2016 while equity market issuances were \$655 billion. Typically, the process of issuing debt begins with the corporation obtaining a necessary credit rating for the issue from a major rating agency, such as Moody's or Standard & Poor's, and ends with an investment banker bringing the issue to market. The rating agency's very existence depends upon being independent, along with the associated credibility the public attributes to the ratings it issues. This logical assumption of independence is corroborated by research indicating ratings are motivated more by reputation protection than rating fees obtained from the issuers (Covitz and Harrison 2003).

The assigned credit rating contains important implications about the bond issue and its subsequent yield. Yield spread between rating categories can be substantial, often resulting in a difference of tens of millions of dollars in interest over the life of the obligation. For example, the yield spread between Baa and Ba rated debt often averages over 100 basis points (Standard and Poor's Credit Week 2010) which can mean a difference in nominal payments of over \$100 million over twenty years for a \$500 million bond issue. There are also many regulatory requirements in the United States and abroad that are specified in terms of a firm's assigned credit rating. Several agencies allow investments to be made only in the top rating categories (e.g. BBB- and above), typically referred to as "Investment Grade" debt. Ratings also play a prominent role in the international rules determining the level of capital banks must maintain. The fact that regulatory agencies define requirements partially based on independent ratings indicates the importance and degree to which the credit rating process is ingrained in the market system.

The economic impact of rating upgrades and downgrades is significant. There is substantial empirical evidence in the finance and accounting literature documenting the importance and information content of credit ratings and changes in credit ratings. Research indicates credit ratings contain pricing-relevant information and this information has measureable effects on the debt value and equity value of the firm (Kliger and Sarig 2000). Research also provides evidence of stock price movement and abnormal returns after credit rating changes (Holthausen and Leftwich 1986; Kim and Nabar 2007; Dichev and Piotroski 2001). In summary, there is substantial empirical evidence indicating the importance of credit ratings to the capital markets. This makes credit ratings an interesting and important area in which to explore the effects of financial statement comparability.

2.2 Cost of Debt

We also examine the effects of comparability on a firm's cost of debt. While credit ratings are determined by a team of rating analysts, a firm's cost of debt is determined more directly by market investors and provides an additional view of the effects of comparability on the overall value of the firm. A significant body of research exists which explores various accounting and finance related questions by investigating effects on a firm's cost of debt capital. For example, Mansi et al. (2011) examine the relation between analyst forecasts and cost of debt and find the impact of analysts is most evident when uncertainty is higher. Anderson et al. (2004) examine board of director characteristics and accounting report integrity and find more independent board of directors and larger board size result in a lower cost of debt. They also find fully independent audit committees that meet frequently appear to increase the quality of accounting reports which correspondingly helps to reduce the cost of debt capital. Mansi et al. (2004) find firms engaging a Big 4 auditor achieve lower debt costs through increasing the perceived reliability of their financial statements, as well as obtaining the additional insurance protection Big 4 auditors are conjectured to provide by ostensibly protecting investors from future losses from audit failure. In a similar vein, Pittman and Fortin (2004) find selecting Big 4 auditors has a positive effect in reducing cost of debt capital for newly public firms.

Other research utilizing a cost of debt framework examines earnings benchmarks, accruals quality, and earnings quality. Jiang (2008) finds the most salient of three common earnings benchmarks in the debt markets is the benchmark related to meeting positive earnings. Francis et al. (2005) conduct extensive analyses of the market pricing of accruals quality and find firms with higher accounting accruals quality (AQ) have lower ratios of interest expense to interest-bearing debt. Francis et al. (2008) examine the relations that exist between voluntary disclosure, earnings quality, and cost of capital and find that earnings quality plays a critical complementary role in evaluating the effects of voluntary disclosure on the cost of debt. Overall, there is considerable literature that illustrates how examination of the cost of debt can provide beneficial insights concerning questions germane to the accounting profession. This literature also provides additional incentive and support for the importance of investigating how comparability might affect the cost of debt capital.

2.3 The Benefit of Financial Statement Comparability

Recently, several studies have examined how capital market participants view financial statement comparability. Comparability in this sense refers to how firms in similar environments incorporate transactions in their published financial statements. These transactions are then reflected in stock returns. The more closely a firm's information/return process matches another represents increased comparability. This basic premise is the reasoning behind the widely used measure developed by De Franco, Kothari, and Verdi (2011). Recognizing that each firm will make different decisions regarding inclusion, exclusion, timing expectations, as well as discount rates that impact the amount of a particular transaction recorded on the financial statements. This process will affect the earnings reported by different firms, and thus, future stock returns will vary.

Most current work examining financial statement comparability is based on De Franco et al. (2011) which develop a mapping of earnings to returns and document that higher comparability is positively related to analyst following and forecast accuracy. Also, the authors find that comparability is negatively related to analyst forecast dispersion. Sohn (2016) finds that firms with greater comparability have lower levels of accrual earnings management and higher levels of real earnings management, indicating that managers are constrained from a financial reporting perspective when they have high comparability. Imhoff, Seavey, and Smith (2017) extend this stream of research by examining the impact comparability has on the cost of equity capital. They find that greater comparability leads to lower cost of equity capital and this effect is most pronounced for firms with high information asymmetry. When examining the acquisitions, Chen, Collins, Kravet, and Mergenthaler (2018) find that as target comparability increases, so does acquisition profitability. Comparability is most beneficial to acquirers when information asymmetry of the target is higher, similar to results reported by Imhoff et al. (2017).

In addition to equity market studies, there has been some examination of the effects of comparability in the debt market setting as well. In the syndicated loan market setting, Fang, Li, Xin, and Zhang (2016) find that comparability impacts debt issue metrics. Fang et al. (2016) find that lower loan spreads and lower likelihood of collateral requirements are related to higher levels of comparability. Additionally, more comparable firms were more likely to have longer maturities and performance pricing, indicating that banks who issue syndicated loans offer more beneficial terms to firms with financial statements that have higher comparability. Kim, Kraft, and Ryan (2013) develop and use a different measure of comparability based on Moody's adjustments to reported earnings. They adjust interest coverage ratio and non-recurring items. Their findings show that firms with more comparability on those measures have seasoned bonds that have lower bid-ask spreads, lower credit spreads, and steeper one- to five-year credit default swap term structure.

Based on the above discussion, we propose and test the following two hypotheses.

H1: Comparability is positively related to initial bond ratings.

H2: Comparability is negatively related to initial bond yields.

3. Variable Definitions and Research Design

3.1 Measure of Comparability

Financial accounting systems can be viewed as to map economic events to accounting numbers on financial statements. We follow De Franco et al. (2011) to empirically estimate accounting comparability as how closely two firms' financial statements reflect similar economic events. DeFranco et al. (2011) use stock returns to represent economic events and earnings to proxy for financial statements. Specifically, we first estimate the following equation over the previous 16-quarter period.

$$Earnings_{it} = \alpha_i + \beta_i Return_{it} + \varepsilon_{it}$$
 (1)

Where Earnings is the ratio of quarterly net income before extraordinary items to the market value of equity at the beginning of the quarter, and return is stock price return during the quarter.

We then estimate the expected earnings of each firm, i and j, as a function of the same economic event,

$$E(Earnings)_{iit} = \hat{\alpha}_i + \hat{\beta}_i Return_{it}$$
 (2)

$$E(Earnings)_{ijt} = \hat{\alpha}_i + \hat{\beta}_i Return_{it}$$
 (3)

where $E(Earnings)_{iit}$ is the predicted earnings of firm i given firm i's stock returns in period t and $E(Earnings)_{ijt}$ is the predicted earnings of firm j given firm i's stock returns in period t. By using firm i's stock return in both predictions, we measure the comparability of mappings between firm i and firm j for the same event (i.e., $Return_{ij}$).

We calculate comparability between firm i and firm j during the 16-quarter estimation period from equations (2) and (3) as the negative value of the average absolute difference between the predicted earnings using firm i's and firm j's earnings functions:

$$Comparability_{ijt} = -\frac{1}{16} \times \sum_{t=15}^{t} \left| E(Earnings_{iit}) - E(Earnings_{ijt}) \right| \tag{4}$$

Greater values indicate higher comparability and the maximum value of the measure is zero. We calculate four different measures of comparability for our empirical analyses. After we estimate equations (2) and (3) for every firm i – firm j combination within the same two-digit Standard Industry Classification (SIC) industry, we rank all the values of $Comparability_{ijt}$ and take average of highest four values. We refer to that measure as $Comp_Four$. Similarly we take the average of highest ten values of $Comparability_{ijt}$ to get our second measure, $Comp_Ten$. Our third and four measures, $Comp_Mean$ and $Comp_Med$, are the mean and median value of all $Comparability_{ijt}$ for all firms j in the same industry as firm i during period t.

3.2 Research Methodology

Our bond rating model is based on the econometrically sound work of Kaplan and Urwitz (1979) which remains robust and continues to be the foundation utilized by many researchers to examine contemporary issues (e.g. Francis et al. 2005; Ashbaugh-Skaife et al. 2006; Francis et al. 2008; Crabtree and Maher 2009). We utilize this basic model and expand it by including our measures of comparability. The basic bond rating model can be represented as follows (all independent variables are measured in the year prior to the bond issuance so that the financial information is available at the bond issuance.):

$$RATING = \beta_0 + \beta_1 COMPARABILITY + \beta_2 SUB + \beta_3 Log ASSETS + \beta_4 LEVERAGE + \beta_5 INCOME + \beta_6 BETA + \beta_7 AA_PM + \sum \beta_i FFIND_i + \sum \beta_t YEAR_t$$
 (5)

where *RATING* = ordinal representation of the issue's initial rating where Aaa = 30, Aa1 = 29, Aa2=28, etc.

COMPARABILITY = measures of accounting comparability described previously. Specifically:

Comp_Four (Ten) = The average of four (ten) firms with the highest comparability for firm i.

Comp_Mean (Med) = Industry mean (median) of comparability for firm i for all firms in firm i's industry.

SUB

= 1 for subordinated bonds and 0 otherwise; the subordination status is stated on the bond indenture and is expected to be associated with higher risk, therefore exhibiting a negative association with bond rating;

Log ASSETS = log of total assets; Log ASSETS proxies for firm size and is expected to have a positive

association with bond rating;

LEVERAGE = long-term debt divided by total assets; LEVERAGE represents the relative amount of debt

currently incurred by the firm; LEVERAGE is expected to have a negative association with

bond rating;

INCOME = operating income for the year divided by total assets; INCOME is expected to have a

positive association with bond rating;

BETA = the firm's common stock Beta for the last 5 years calculated using monthly returns; firms

with a higher beta are considered riskier, hence BETA is expected to have negative

association with bond rating;

AA_PM = the performance matched discretionary accruals

FFIND = Fama French 48-industry classifications;

YEAR = Year dummy variable

To examine the effect of comparability on bond pricing, it is necessary to utilize a two-stage regression procedure following Anderson et al. (2004), Mansi et al. (2004), and Mansi et al. (2011). Two-stage regression is appropriate when examining bond yields because the bond rating has already been assigned to the issue prior to its sale to the public. To disentangle the effects of the rating and the yield, we estimate Equation 5 in the first stage, and then pass the residual onto our bond yield model to represent the orthogonal rating information (ORTHRATE). Consistent with prior research (e.g. Anderson et al. 2004; Mansi et al. 2011), we define bond yield as the spread between the issue yield and its duration-matched Treasury equivalent.

Our yield model can be represented as:

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\begin{split} \text{YLDSPREAD} &= \beta_0 + \beta_1 \text{ COMPARABILITY} + \beta_2 \text{ ORTHRATE} + \beta_3 \text{ SIZE} + \beta_4 \text{ CALL} + \beta_5 \text{ INTVOL} + \beta_6 \text{ SUB} + \beta_7 \\ & \text{Log ASSETS} + \beta_8 \text{ LEVERAGE} + \beta_9 \text{ INCOME} + \beta_{10} \text{ BETA} + \beta_{11} \text{ AA\_PM} + \sum \beta_i \text{ FFIND}_i + \sum \beta_t \\ & \text{YEAR}_t + \epsilon \end{split}
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Where:

YLDSPREAD = the spread between the issue yield and its duration-matched Treasury equivalent;

COMPARABILITY= measures of accounting comparability, as defined above for Eq. (5).

ORTHRATE = the security specific adjusted orthogonal credit rating which is the residual from the first

stage regression, Equation 5.

SIZE = the natural log of the bond issue amount;

CALL = the ratio of years to call divided by years to maturity;

INTVOL = volatility of the prior ten days of the duration-matched Treasury equivalent;

The remaining independent variables are as defined earlier for Equation 5.

4. Data and Results

4.1 Sample Selection

We examine new corporate bond issues rated by Moody's bond rating agency and issued during the time period January 1990 to December 2014. The sample consists of non-financial firms that issue bonds backed solely by the issuer's ability to pay. No convertible bonds, mortgage bonds, asset-backed bonds, or deferred interest bonds are included. This allows us to examine bonds which are based entirely on the issuing firm's default risk and not the risk of another underlying asset or option. The sample is restricted to new issues because this helps ensure the rating is the result of a detailed and complete recent analysis conducted by bond rating analysts and avoids problems related to any potential delay in changing existing ratings which has been expressed in the popular press (e.g., Wall Street Journal 2011, 2012). Moreover, because each issue is new, we do not have to worry about any timing delays in the reflection of current price due to bonds that are thinly traded. The initial sample consists of 5,648 new debt issues from firms listed on Compustat. For companies with more than one debt issues, we combine them as one issue which eliminate 2,554 observations. There are 1,113 new issues that do not have the necessary variables on Compustat and CRSP to calculate the accounting comparability measures and other control variables. This results in a sample of 1,981 unique firm-year observations for 587 firms during the period 1990 through 2014.

Table 1 reports the sample distribution. The sample is distributed evenly over the years. There is no evidence of industry clustering in the sample. The industries that contain the highest percentage of issuing new debts include: food, tobacco, textiles, paper, and chemicals (33.27%), manufacturing, machinery and electronics (30.49%), and agriculture, mining, oil, and construction (12.37%)

Table 1. Sample Distribution

Panel A. Sample distribution - Year

	Frequency	Percent
1990	30	1.51
1991	76	3.84
1992	73	3.69
1993	80	4.04
1994	36	1.82
1995	63	3.18
1996	69	3.48
1997	92	4.64
1998	136	6.87
1999	110	5.55
2000	45	2.27
2001	108	5.45
2002	80	4.04
2003	60	3.03
2004	40	2.02
2005	82	4.14
2006	81	4.09
2007	60	3.03
2008	61	3.08
2009	97	4.90
2010	104	5.25
2011	82	4.14
2012	110	5.55
2013	108	5.45
2014	98	4.94
Total	1,981	100.00

Panel B. Sample distribution - Industry

Industry (1-digit SIC)	Frequency	Percent
0-1 Agriculture, mining, oil, and construction	245	12.37%
2 Food, tobacco, textiles, paper, and chemicals	659	33.27%
3 Manufacturing, machinery, and electronics	604	30.49%
4 Transportation and communications	150	7.57%
5 Wholesale and retail	166	8.38%
7 Services	119	6.01%
8-9 Health, legal, and educational services and other	38	1.92%
Total	1981	100.00%

Table 2 reports the descriptive statistics for our sample. The mean yield spread is 184 basis points, with a median of 134 basis points and a standard deviation of 149 basis points. The average firm in the sample has Moody's credit rating of Baa1, a leverage ratio of 25%. The firms are profitable, with ROA of 7%, and relatively large, with mean total assets of \$6,634 million. Only 7 percent of the sample issued subordinated bonds.

In Table 2 we also present the descriptive statistics of the four comparability measures. Recall that all the measures are negative by construction and a value closer to zero denotes more comparability. *Comp_Four* and *Comp_Ten* have means values of -0.40 and -0.62 respectively while *Comp_Mean* and *Comp_Med* have mean values of -2.83 and -1.85 respectively. All values are similar to the values reported by prior studies.

Table 2. Descriptive Statistics

Variable	N	Mean	Std Dev	Q1	Median	Q3
Rating	1981	22.47	3.52	21.00	23.00	25.00
Yield Spread	1981	1.84	1.49	0.78	1.34	2.45
Comp_Four	1981	-0.40	0.68	-0.37	-0.17	-0.09
Comp_Ten	1981	-0.62	0.96	-0.62	-0.28	-0.15
Comp_Mean	1981	-2.83	1.80	-3.37	-2.46	-1.84
Comp_Med	1981	-1.85	1.85	-2.11	-1.36	-0.98
SUB	1981	0.07	0.26	0.00	0.00	0.00
LOGASSET	1981	8.80	1.39	7.86	8.83	9.71
LEVERAGE	1981	0.25	0.14	0.15	0.23	0.33
INCOME	1981	0.07	0.07	0.03	0.07	0.11
BETA	1981	1.06	0.46	0.75	1.02	1.32
AA_PM	1981	0.03	0.07	-0.01	0.02	0.07

This table provides descriptive statistics for our sample. Variables are defined as follows. *RATING* is the ordinal representation of the issue's initial rating where Aaa equals 30 and decreases for each rating category below; *Yield Spread* is the spread between the issue yield and its duration-matched Treasury equivalent; *Comp_Four (Ten)* is a measure of financial statement comparability based on how closely economic events are represented in earnings across firms within an industry. *Comp_Four (Ten)* is the average of four (ten) firms with the highest comparability for firm i. *Comp_Mean (Med)* is the industry mean (median) of comparability for firm i for all firms in firm i's industry. *LEVERAGE* is the firm's long-term debt scaled by total assets; *SUB* is an indicator variable that equals 1 if the issue is subordinated, and 0 otherwise; *LOGASSETS* is natural logarithm of the firms total assets; *INCOME* is the operating income of the firm scaled by total assets; *BETA* is the firm's common stock Beta; firms with a higher beta are considered riskier, hence *BETA* is expected to have negative association with bond rating; *AA_PM* is the performance matched discretionary accruals

Table 3 provides correlation coefficients between bond credit rating, yield spread, comparability measures, and independent variables. Three observations are noteworthy. First, all four comparability measures are highly correlated, with correlation coefficients from 0.70 to 0.96, indicating that there is no construct validity issue. That is the four measures proxy for the same construct – accounting comparability. Second, credit ratings are positively related to each of the four comparability measures. Thus, from a univariate perspective, high accounting comparability is associated with a high credit rating. Third, in a similar manner, each of the four comparability measures is negatively associated with yield spread. This suggests, if a company accounting is more comparable to other companies in the same industry, the bond investors require a lower premium for the bonds. Taken together, the correlation table provides initial evidence that bond rating agencies and bond investors view accounting comparability positively and adjust the bond rating and pricing accordingly.

Table 3. Correlation matrix

		1	2	3	4	5	6	7	8	9	10	11
1	Rating	1										
2	Yield Spread	-0.73	1									
3	Comp_Four	0.40	-0.38	1								
4	Comp_Ten	0.40	-0.37	0.96	1							
5	Comp_Mean	0.29	-0.28	0.71	0.70	1						
6	Comp_Med	0.33	-0.31	0.78	0.76	0.95	1					
7	SUB	-0.48	0.36	-0.14	-0.13	-0.09	-0.11	1				
8	LOGASSET	0.50	-0.31	0.10	0.10	-0.03	0.04	-0.33	1			
9	LEVERAGE	-0.52	0.40	-0.24	-0.23	-0.16	-0.21	0.31	-0.24	1		
10	INCOME	0.31	-0.29	0.21	0.19	0.02	0.11	-0.18	0.17	-0.26	1	
11	BETAV	-0.20	0.20	-0.21	-0.21	-0.21	-0.22	0.02	0.04	0.01	-0.06	1
12	AA_PM	0.01	0.03	0.05	0.09	-0.20	-0.14	-0.04	0.15	-0.03	0.22	0.04

This table reports Pearson correlation for bond rating, yield spread, comparability measures and other control variables. *RATING* is the ordinal representation of the bond issue's initial rating. *YLDSPREAD* is the spread between the issue yield and its duration-matched Treasury equivalent; for firms with multiple bond issues, a weighted average spread is calculated based on the bond's issue size. *Comp_Four (Ten, Mean, Med)* are the four measures of financial statement comparability. *Log ASSETS* represents the natural log of total assets. *LEVERAGE* is long-term debt divided by total assets. *INCOME* is operating income for the year divided by total assets. *AA_PM* is performance matched discretionary accruals. Bold indicates significance at a p-value < 0.01.

4.2 Bond Ratings and Comparability

To empirically test whether bond rating agencies incorporate comparability into their ratings, we employ ordinal logistic regression with results reported in Table 4. The summary statistics indicate the models are robust (Pseudo R^2 s over 72%), suggesting the variables explain a significant amount of variation of bond ratings. All control variables are significant and have their correct theoretical signs.

Our variables of interest are the measures of comparability. We find that the coefficients of all four comparability measures are significantly positive. For example, the coefficient on $Comp_Four$ is 0.853 with p-value < 0.01, suggesting that for firms with higher accounting comparability, they are more likely to receive a higher bond rating supporting hypothesis H1. The results from the other three measures provide consistent supporting evidence for hypothesis H1.

Table 4. Logistic Regression Results - Bond Rating Models (n=1,981)

 $RATING = \beta_0 + \beta_1 \ COMPARABILITY + \beta_2 \ SUB + \beta_3 \ Log \ ASSETS + \beta_4 \ LEVERAGE + \beta_5 \ INCOME + \beta_6 \ BETA \\ \beta_7 \ AA_PM + \sum \beta_i \ FFIND_i + \sum \beta_t \ YEAR_t$

	Comp_Fo	Comp_Four				Comp_Ten			
	Estimate	Wald Chi-Sq		OR	Estimate	Wald Chi-Sq		OR	
Comparability	0.853	142.00	***	2.35	0.590	135.60	***	1.80	
SUB	-2.333	138.06	***	0.10	-2.338	138.76	***	0.10	
LOGASSET	1.096	683.52	***	2.99	1.089	676.10	***	2.97	
LEVERAGE	-6.257	295.15	***	0.00	-6.324	301.75	***	0.00	
INCOME	13.231	259.62	***	>999	13.406	267.07	***	>999	
BETAV	-0.646	40.41	***	0.52	-0.645	40.18	***	0.53	
AA_PM	-3.481	23.24	***	0.03	-3.753	26.96	***	0.02	
Industry FE	Yes				Yes				
Year FE	Yes				Yes				
Loglikelihood	2597.25				2590.33				
Pseduo R ²	0.7336				0.7327				

	Comp_Mean				Comp_Med			
		Wald			Wald			
	Estimate	Chi-Sq		OR	Estimate	Chi-Sq		OR
Comparability	0.312	106.90	***	1.37	0.307	114.16	***	1.36
SUB	-2.326	137.22	***	0.10	-2.367	141.06	***	0.09
LOGASSET	1.096	684.60	***	2.99	1.090	677.09	***	2.97
LEVERAGE	-6.326	302.05	***	0.00	-6.205	289.76	***	0.00
INCOME	13.884	288.07	***	>999	13.458	269.63	***	>999
BETAV	-0.659	41.93	***	0.52	-0.639	39.30	***	0.53
AA_PM	-2.737	14.35	***	0.07	-2.838	15.45	***	0.06
Industry FE	Yes				Yes			
Year FE	Yes				Yes			
Loglikelihood	2562.54				2571.19			
Pseduo R ²	0.7289				0.7301			

This table provides the results on the relation between bond rating and comparability and other control variables. Industry and year indicator variables are included in the regression but the coefficient estimates and statistics are suppressed, p-values are based on two-tailed tests. RATING is the ordinal representation of the issue's initial rating. Comp_Four (Ten, Mean, Med) are the four measures of financial statement comparability. SUB = 1 for subordinated bonds and 0 otherwise; the subordination status is stated on the bond indenture and is expected to be associated with a higher risk, therefore exhibiting a negative association with bond rating. Log ASSETS represents the log of total assets; Log ASSETS proxies for firm size and is expected to have a positive association with bond rating. LEVERAGE is firm long-term debt divided by total assets; LEVERAGE represents the relative amount of debt currently incurred by the firm; LEVERAGE is expected to have a negative association with bond rating. INCOME represents the

operating income for the year divided by total assets; *INCOME* is expected to have a positive association with bond rating. *BETA* is the firm's common stock Beta; firms with a higher beta are considered riskier, hence *BETA* is expected to have negative association with bond rating. *AA PM* – performance matched discretionary accruals

4.3 Bond Yield Spread and Comparability

An alternative approach to investigate the effect of accounting comparability on cost of debt is to examine the impact of comparability on bond yield spreads. While bond ratings reflect the decision of rating analysts, the yield at issuance is a collective market measure of creditors' confidence in the default risk of the issue. By examining new issues, we are able to observe a concise assessment by creditors of the issue without having to worry about timing, thin trading, and other informational delays. Table 5 provides the results from the regressions.

Table 5. Regression Results - Bond Yield Models (n=1,981)

$$\begin{split} \text{YLDSPREAD} &= \beta_0 + \beta_1 \text{ COMPARABILITY} + \beta_2 \text{ ORTHRATE} + \beta_3 \text{ SIZE} + \beta_4 \text{ CALL} + \beta_5 \text{ INTVOL} + \beta_6 \text{ SUB} + \beta_7 \\ & \text{Log ASSETS} + \beta_8 \text{ LEVERAGE} + \beta_9 \text{ INCOME} \\ & + \beta_{10} \text{ BETA} \\ & + \beta_{11} \text{ AA_PM} + \sum \beta_i \text{ FFIND}_i + \sum \beta_t \text{ YEAR}_t + \epsilon \end{split}$$

	Comp_Fo	ur		Comp_Ten			
Parameter	Estimate	t Value		Estimate	t Value		
Intercept	4.552	12.66	***	4.464	12.39	***	
COMPARABILITY	-0.392	-7.98	***	-0.271	-7.89	***	
RATERES	-0.255	-16.84	***	-0.256	-16.91	***	
SIZE	0.017	0.63		0.016	0.61		
CALL	-0.026	-0.60		-0.029	-0.66		
INTVOL	0.394	1.73	*	0.361	1.59		
SUB	0.945	8.42	***	0.954	8.49	***	
LOGASSET	-0.336	-13.80	***	-0.334	-13.66	***	
LEVERAGE	1.884	8.54	***	1.932	8.79	***	
INCOME	-5.359	-11.29	***	-5.429	-11.50	***	
BETA	0.330	5.37	***	0.329	5.33	***	
AA_PM	0.398	0.97		0.511	1.24		
Industry FE	Yes			Yes			
Year FE	Yes			Yes			
Adj. R ²	0.7336			0.689			

	Comp_M	ean		Comp_Med			
Parameter	Estimate	t Value		Estimate	t Value		
Intercept	4.478	12.41	***	4.601	12.80	***	
COMPARABILITY	-0.131	-5.45	***	-0.118	-5.35	***	
RATERES	-0.261	-17.58	***	-0.261	-17.60	***	
SIZE	0.013	0.48		0.013	0.47		
CALL	-0.028	-0.64		-0.026	-0.58		
INTVOL	0.367	1.60		0.380	1.67	*	
SUB	0.967	8.57	***	0.972	8.59	***	
LOGASSET	-0.339	-13.87	***	-0.336	-13.70	***	
LEVERAGE	1.950	8.90	***	1.902	8.53	***	
INCOME	-5.787	-12.09	***	-5.645	-11.70	***	
BETA	0.345	5.69	***	0.338	5.50	***	
AA_PM	0.118	0.29		0.175	0.43		
Industry FE	Yes			Yes			
Year FE	Yes			Yes			
Adj. R ²	0.6872			0.6863			

Standard errors are clustered by firm. This table provides the results on the relation between bond yield spread and comparablity and other control variables. A two stage regression procedure is utilized following prior research, p-values are based on two-tailed tests. *YLDSPREAD* is the spread between the issue yield and its duration-matched Treasury equivalent; for firms with multiple bond issues, *YLDSPREAD* is calculated as a weighted average based on the bond's issue size. *Comp_Four* (*Ten, Mean, Med*) are the four measures of financial statement comparability. *ORTHRATE* is the security specific adjusted credit rating; it is the residual from the first stage regression. ORTHRATE = RATING – ($\beta_0 + \beta_1 EM + \beta_2 SUB + \beta_3 ASSETS + \beta_4 DEBT + \beta_5 INCOME + \beta_6 BETA$). *SIZE* is the natural log of the bond issue amount. *CALL* is the ratio of years to call divided by years to maturity. *INTVOL* is the volatility of the prior ten days of the duration-matched Treasury equivalent. *SUB* = 1 for subordinated bonds and 0 otherwise. *Log ASSETS* = the log of the firms total assets. *LEVERAGE* = the firm's long-term debt scaled by total assets. *INCOME*= the operating income of the firm scaled by total assets. *BETA* = the firm's common stock beta. *AA_PM* – performance matched discretionary accruals

Our results for bond yields are consistent with those presented previously for ratings. All four comparability measures are negatively related to lower bond yields at (p-value<0.01), providing support for H2. These are consistent with investors requiring an lower yield premium to purchase bonds from firms with high comparability.

5. Conclusion

The FASB conceptual framework defines comparability as "the qualitative characteristic that enables users to identify and understand similarities in, and differences among, financial statement items." The accounting comparability is viewed as an important accounting quality (Kothari et al. 2010). Therefore, we examine the impact of accounting comparability on cost of new debt issues. New debt issues provide us an ideal sample to investigate this question in that the ratings are provided by the independent bond rating analysts after a detailed and complete analysis. It also avoids any timing delays in the reflection of current price duo to bonds that are thinly traded.

Our empirical results indicate that firms with higher accounting comparability receiver higher bond ratings. We also find that bond yields are lower for firms with higher accounting comparability, consist with that bond investors require lower bond yield premium when firm's accounting comparability is high. We make contribution to the general research literature on the effect of financial reporting quality on capital market participants, specifically to the benefit of financial statement comparability. Our findings are of interest to regulators, managers and analysts.

References

- Anderson, R.C., S. A. Mansi, D. M. Reeb. (2004). Board characteristics, accounting report integrity, and the cost of debt. *Journal of Accounting and Economics*, 37(3), 315-342. https://doi.org/10.1016/j.jacceco.2004.01.004
- Ashbaugh-Skaife, H., D. Collins, R. LaFond. (2006). The Effects of Corporate Governance on Firms' Credit Ratings. *Journal of Accounting and Economics*, 42, 203-243. https://doi.org/10.1016/j.jacceco.2006.02.003
- Bharath, S. T., S. Dahiya, A. Saunders, & A. Srinivasan. (2011). Lending relationships and loan contract terms. *Review of Financial Studies*, 24(4), 1141–1203. https://doi.org/10.1093/rfs/hhp064
- Chen, C., Collins, D. W., Kravet, T., & Mergenthaler, R. (2018). Financial Statement Comparability and the Efficiency of Acquisition Decisions. *Contemporary Accounting Research*, 35(1), 164-202. https://doi.org/10.1111/1911-3846.12380
- Covitz, D.M. & P. Harrison. (2003). Testing conflicts of interest at bond rating agencies with market anticipation: Evidence that reputation incentives dominate. Working paper, Federal Reserve Board.
- Crabtree, A.D., & J.J. Maher. (2009). The influences of differences in taxable income and book income on the bond credit market. *Journal of the American Taxation Association*, 31(1), 75-99. https://doi.org/10.2308/jata.2009.31.1.75
- De Franco, G., S. P. Kothari, & R. S. Verdi. (2011). The benefits of financial statement comparability. *Journal of Accounting Research*, 49(4), 895–931. https://doi.org/10.1111/j.1475-679X.2011.00415.x
- Dichev, I. D., & Piotroski, J. D. (2001). The long run stock returns following bond ratings changes. *Journal of Finance*, 56(1), 173-203. https://doi.org/10.1111/0022-1082.00322
- Easley, D., & M. O'Hara. (2004). Information and the cost of capital. *Journal of Finance*, 59(4), 1553–1583. https://doi.org/10.1111/j.1540-6261.2004.00672.x
- Fang, X., Li, Y., Xin, B., & Zhang, W. (2016). Financial statement comparability and debt contracting: Evidence from the syndicated loan market. *Accounting Horizons*, 30(2), 277–303. https://doi.org/10.2308/acch-51437
- Francis, J., R. LaFond, P. Olsson, & K. Schipper. (2005). The market pricing of accruals quality. *Journal of Accounting and Economics*, 39(2), 295-327. https://doi.org/10.1016/j.jacceco.2004.06.003
- Francis, J., D. Nanda, & P. Olsson. (2008). Voluntary disclosure, earnings quality, and cost of capital. *Journal of Accounting Research*, 46(1), 53-99. https://doi.org/10.1111/j.1475-679X.2008.00267.x
- Holthausen, R. W., & Leftwich, R. W. (1986). The effect of bond rating changes on common stock prices. *Journal of Financial Economics*, 17, 57-90. https://doi.org/10.1016/0304-405X(86)90006-1
- Imhof, M. J., Seavey, S. E., & Smith, D. B. (2017). Comparability and cost of equity capital. *Accounting Horizons*, 31(2), 125-138. https://doi.org/10.2308/acch-51710
- Jiang, J. (2008). Beating earnings benchmarks and the cost of debt. *The Accounting Review*, 83(2), 377-416. https://doi.org/10.2308/accr.2008.83.2.377
- Kim, S., P. Kraft, & S. Ryan. (2013). Financial statement comparability and credit risk. *Review of Accounting Studies*, 18(3), 783–823. https://doi.org/10.1007/s11142-013-9233-z
- Kaplan, R.S. & G. Urwitz. (1979). Statistical models of bond ratings. *Journal of Business*, 52, 231-262. https://doi.org/10.1086/296045
- Kim, Y. & S. Nabar. (2007). Bankruptcy probability changes and the differential informativeness of bond upgrades and downgrades. *Journal of Banking and Finance*, *31*(12), 3843-3861. https://doi.org/10.1016/j.jbankfin.2007.02.008
- Lambert, R., Leuz, C., & Verrecchia, R. E. (2007). Accounting information, disclosure, and the cost of capital. *Journal of Accounting Research*, 45(2), 385–420. https://doi.org/10.1111/j.1475-679X.2007.00238.x
- Kliger, D. & O. Sarig. (2000). The information value of bond ratings. *Journal of Finance*, 55(6), 2879-2902. https://doi.org/10.1111/0022-1082.00311
- Kothari, S.P., Karthik Ramanna & Douglas J. Skinner. (2010). Implications for GAAP from an Analysis of Positive Research in Accounting. *Journal of Accounting & Economics*, 50(2-3), 246-286. https://doi.org/10.1016/j.jacceco.2010.09.003

- Mansi, S. A., W. F. Maxwell, & D. P. Miller. (2004). Does auditor quality and tenure matter to investors? Evidence from the bond market. *Journal of Accounting Research*, 42(4), 755-793. https://doi.org/10.1111/j.1475-679X.2004.00156.x
- Mansi, S. A., W. F. Maxwell, & D. P. Miller. (2011). Analyst forecast characteristics and the cost of debt. *Review of Accounting Studies*, 16 (March), 116-142. https://doi.org/10.1007/s11142-010-9127-2
- Pittman, J.A., & S. Fortin. (2004). Auditor choice and the cost of debt capital for newly public firms. *Journal of Accounting and Economics*, 37(1), 113-136. https://doi.org/10.1016/j.jacceco.2003.06.005
- Sohn, B C. (2016). The effect of accounting comparability on accrual-based and real earnings management. *Journal of Accounting and Public Policy*, *35*, 513-539. https://doi.org/10.1016/j.jaccpubpol.2016.06.003
- Thomson Reuters debt http://dmi.thomsonreuters.com/Content/Files/2016%20Debt%20Capital%20Markets%20Review.pdf
- Thomson Reuters equity http://share.thomsonreuters.com/general/PR/ECM_4Q_2016_E.pdf
- Wall Street Journal. (2001). Why credit agencies didn't switch off Enron. Wall Street Journal November 29: C1.
- Wall Street Journal. (2002). Moody's and S&P may accelerate process for cutting credit ratings. Wall Street Journal January 22: C1.